

## POVERTY INCIDENCE IN NIGERIA: AN APPLICATION OF THE BINOMIAL TEST

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### ABSTRACT

*This study considers an alternative estimation of the incidence of poverty based on the binomial test. The core objective is to explore an available alternative in the presence of limitations posed by consumption measurement problems, inadequate technical capacity and paucity of funds to conduct reliable household surveys. It is believed that the alternative approach considered will be less demanding and cost effective. This is the Nigerian situation with dwindling resources for the collection of necessary household survey data. The incidence of poverty is estimated from the binomial test using a recent general household survey panel data with per capita expenditure as the proxy for poverty. The estimate of the incidence of poverty from the binomial test compared well with the estimate expected conventionally. The alternative estimation procedure from the binomial test may therefore be adopted by socio and economic researchers especially in less developed nations because it gives an indication of the increase or decrease in poverty incidence. This will indeed aid assessment of poverty intervention programmes.*

**Keywords:** Incidence of poverty, Binomial test, Index of poverty, Poverty measurement, Multidimensional poverty, Socio and Economic researchers

### INTRODUCTION

Poverty in Nigeria as noted by Anyawu (2012) is on the increase. He observed that the main reason for this is bad implementation of policies by government. These policies, according to him failed to concentrate on those who are poor. The level of poverty incidence in Nigeria in 2010 was 69% (Kale, 2012). He noted again that both the North East and North West geopolitical zones had the highest poverty incidences of 76.3% and 77.7%, respectively, while the South West zone had the lowest incidence of 59.1%. Furthermore,

according to Kale (2012), the Sokoto state had the highest incidence of 86.4%, while Niger state had the least incidence of 43.6%.

Choosing an appropriate index of poverty from available options is cardinal in the course of poverty measurement and this is dependent on an appropriate identification of the poor in a given population (Fajardo-Duka, 1992)

The definition of who the poor are is not an easy matter. This is because poverty is seen from different angles by different people. This perhaps confirms the multi-dimensional nature of poverty. A household might be income poor, consumption poor, asset poor or it might be derived in its access to basic social amenities. All these represent different shades of poverty that a given household might be experiencing at one time or the other. The uni-dimensional approach to poverty is when only income or expenditure is used as a proxy for household poverty. Several authors in Nigeria (Aigbokhan (1997); Canagarajah & Thomas (2001); Ajakaiye & Adeyeye (2001); Akintola & Yusuf (2001); Osinubi (2003); Oyekale et al (2004); Olaniyan & Awoyemi (2006); Adesanoye & Okunmadewa (2007); Olubanjo et al (2007); Omonona et al (2008); Ibrahim & Umar (2008); Oshewolo (2010); Obayelu & Awoyemi (2010); Anyanwu (2011, 2013); Adawo (2011); Olorunsanya & Omotesho (2011); Asogwa et al (2012); Akerele et al (2012); and Okoroafor & Nwaeze (2013)) have adopted this approach in their respective studies on poverty in Nigeria including its various subdivisions. The global adoption of this approach may not be unconnected with its computational appeal as introduced by Foster et al (1984). The approach generally includes the estimation of poverty incidence, depth and severity; the three basic dimensions of poverty.

The concept of multidimensional poverty derived its root from the pioneering work of Amartya Sen (Sen (1979, 1985, 1987). Authors who have also contributed to the growth of the concept include Tsui (2002), Atkinson (2003), Bourguignon and Chakravarty (2003), Alkire and Foster (2007, 2011), Batana (2008), Alkire and Santos (2010), Oyekale (2011), and Levine et al (2014). The main difference between this concept and unidimensional method is its ability to view poverty over a wider spectrum. Thus, it offers a broader scope than the uni-dimensional method.

The motivation for this study is hinged on Sahn and Stifel (2003) who noted that problems of measurement generally limit the use of expenditure measures in less developed nations. They also observed that the quality of expenditure and consumption surveys from these nations are low, coupled with lack of adequate 'technical capacity' on the part of parastatals of government saddled with the responsibility for conducting the surveys; as well as the reduction in foreign aids from international donor agencies to reduce the financial burden of conducting such surveys. It is in the light of this that this present study is being undertaken. This study is an attempt to estimate the poverty incidence from the binomial test in line with the suggestion of Sahn and Stifel (2003, p.465) who opined that "more rapid, less costly, and less demanding approaches for measuring poverty" should be sought. This study is limited to the head count index because it measures the incidence of poverty and

according to Fajardo-Duka (1992, p.162) “it is the most commonly used and easiest to compute”

## 1. Methodology

### 1.1 Traditional Unidimensional Approach

The conventional approach following the uni-dimensional ground breaking work of Foster et al (1984) is the estimation of the three basic dimensions of poverty. These dimensions are poverty incidence, depth and severity. These dimensions are summarized by Foster-Greer-Thorbecke (FGT) generalized poverty index

$$P_{\alpha} = \frac{1}{N} \sum_{i=1}^N \left( \frac{z-y_i}{z} \right)^{\alpha} I(y \leq z) \quad (1)$$

where  $\alpha$  is the aversion index assuming values 0, 1 and 2. The generalized index becomes head count, poverty gap and square poverty gap indices when  $\alpha = 0, 1$  and 2 in that order.  $N$  is the total number of households,  $z$  is the poverty line,  $y_i$  is the per capita household expenditure for the  $i^{\text{th}}$  household and  $I(y \leq z)$  is an indication function assuming value of 0 when  $y > z$  and value of 1 when  $y \leq z$ . The poverty line is a predetermined threshold for classifying any household into poor and non poor. Poor households are those whose per capita expenditure are less than or equal to the poverty line, while non poor households are those whose per capita expenditure are greater than the poverty line. The poverty line is generally defined as

$$z_1 = \frac{2}{3} \text{ mean per capita expenditure} \quad (2)$$

This is the traditional poverty line. This poverty line will be considered in this study with its selected variants

$$z_2 = \text{mean per capita expenditure} \quad (3)$$

$$z_3 = \text{median per capita expenditure} \quad (4)$$

The reason for choosing these poverty lines is based on the comment of Srinivasan (2000) who observed that measures of poverty are sensitive to the poverty lines used in their computations. Additional objective of this study is to see how the hypotheses of the binomial test will be affected by these poverty lines.

As earlier said, this present study is limited to the computation of the incidence of poverty via the head count index. The head count index is defined as

$$P_0 = \frac{1}{N} \sum_{i=1}^N \left( \frac{z-y_i}{z} \right)^0 I(y \leq z) = \frac{q}{N} \quad (5)$$

where  $q$  is the number of poor households in the study population.

## 1.2 The Binomial Test

The test consists of the outcomes of  $n$  independent trials where each outcome falls into two distinct classes; but not both. The classes are generally represented as class1 and class 2 respectively. We shall let the number of observations in the two classes be  $O_1$  and  $O_2$  respectively with  $n$  (total number of observations) equal to the sum of  $O_1$  and  $O_2$ ; that is  $n = O_1 + O_2$ . The two basic assumptions of the test are that each trial has probability ‘ $p$ ’ of belonging to any of the two classes and that the trials are mutually independent. It should be noted that ‘ $p$ ’ is constant over all the trials. The total number of independent trials is generally denoted as  $n$ . This test is a Non Parametric test and specifically suitable for the present study. This is because the household data to be used for analytical considerations allow the use of ordinal scale of measurement. This is true since households are classified as ‘poor’ and ‘non poor’ respectively. Furthermore this test will be considered appropriate since no distribution function is assumed for the per capita expenditure data; a proxy for poverty.

## 1.3 Hypotheses under the Binomial Test

There are basically three forms of hypotheses. These are

$$\begin{array}{ll}
 i. H_0 : p = p^* & ii. H_0 : p = p^* \\
 H_1 : p \neq p^* \text{ (two tailed test)} & H_1 : p > p^* \text{ (one tailed test)} \\
 iii. H_0 : p = p^* & \\
 H_1 : p < p^* \text{ (one tailed test)} & 
 \end{array}$$

Where  $p^*$  is some specified constant such that  $0 \leq p^* \leq 1$ . For (i)  $H_0$  will be rejected if  $T > t_2$  or  $T \leq t_1$ . Also  $H_0$  will be rejected if  $T > t$  and  $T \leq t$  for (ii) and (iii) respectively.  $T$  is the test statistic which refers to the number of observations belonging to class1 ( $O_1$ ) = number of poor households. This implies that  $T = O_1$ . This is so because the binomial test is concerned only with the probability of the outcome ‘‘class1’’. More specifically, we note the following for the tests above. For (i) the level of significance  $\alpha$  corresponds to the two tails of the

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binomial distribution with parameters  $n$  and  $p^*$ , where the size of the upper tail is  $\alpha_2$  and that of the lower tail is  $\alpha_1$  so that  $\alpha_1 + \alpha_2 = \alpha$ . The implication here is that,  $t_1$  and  $t_2$  are obtained from the binomial distribution table such that  $P(Y \leq t_1) = \alpha_1$  and  $P(Y > t_2) = \alpha_2$  where  $Y$  follows the binomial distribution with parameters  $n$  and  $p^*$ . For (ii)  $P(Y > t) = \alpha$  or  $P(Y \leq t) = 1 - \alpha$  and for (iii)  $P(Y \leq t) = \alpha$ . The attractiveness of the binomial test is that it allows us to test hypothesis on whether the poverty incidence has increased or decreased. As earlier noted by Kale (2012), our interest shall be to confirm whether the incidence of poverty has decreased or increased from 69%. In line with the three forms of the binomial test's hypotheses, the following hypotheses shall be of interest:

$$\begin{aligned}
 & H_0 : p = 0.69 \\
 a. & H_0 : p < 0.69 \\
 & H_0 : p = 0.69 \\
 b. & H_0 : p > 0.69
 \end{aligned} \tag{6}$$

**1.4 Large Sample Approximation for the Binomial Test**

When  $n$  is large ( $n \geq 30$ ), the large sample approximation is given below. Let

$$Z^* = \frac{T - E_0(T)}{[Var_0(T)]^{\frac{1}{2}}} = \frac{T - np_0}{\sqrt{np_0q_0}} \sim N(0, 1) \tag{7}$$

where  $np_0$  and  $np_0q_0$  are the expectation and variance of  $T$  under the null hypothesis. From (5), the number of poor households ( $q$ ) is determined from the binomial test by noting that the test statistic  $T$  is also the number of poor households. The implication of this is that the estimate of the head count index from the binomial test is

$$P_0 = \frac{T}{N} \tag{8}$$

This is not very different from the estimate of the head count index given by Kakwani (1993). The decision rules for  $Z^*$  in (7) for the hypotheses mentioned in (6) are to reject the null hypotheses when  $Z^* < Z_\alpha$  (or  $-Z^* > -Z_\alpha$ ) and  $Z^* > Z_\alpha$  (or  $-Z^* < -Z_\alpha$ ) respectively.

## **2. Empirical Illustration**

### **2.1 Description of Data**

The methods presented in this study will be applied to the 2012/2013 Nigeria General Household Survey Panel data. The General Household Panel Survey (GHPS) was conducted by the National Bureau of Statistics (NBS) of Nigeria in 2012 in partnership with the Federal Ministry of Agriculture and Rural Development, the National Food Reserve Agency, the Bill and Melinda Gates Foundation and the World Bank. The 2012/2013 GHPS was the first of its kind in Nigeria. The GHPS was a cross sectional survey of 22,000 households carried out periodically throughout the country. The survey involved the selection of a sub-sample of 5,000 households which became the sample for the GHPS. The information collected on these households included information on agricultural, household income, expenditure and consumption activities. The sample design used in the GHPS was a two stage probability proportional to size sampling scheme. The first stage was the selection of the primary sampling units (PSUs). These were the enumeration areas (EAs). A total of five hundred (500) EAs were selected as PSUs. The second stage was the selection of households. The systematic sampling selection method was used for the selection of ten (10) households from each EA. The selection of ten households per EA was informed from previous general household surveys (GHSs) where ten households per EA were considered adequate. Five hundred (500) EAs were canvassed and 5,000 households were interviewed such that different states had different sample sizes. However, the households considered were not selected with replacement and this made the total number of households interviewed lower than the expected 5,000 households. The choice of the GHPS for a study of this type is ideal since it provides a framework to monitor similar households over time. It should be noted again for emphasis, that the adopted proxy of poverty in this study, is the per capita household expenditure. This is the ratio of total expenditure to the household size for a given household. The poverty lines used are defined in (2), (3) and (4) respectively.

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**2.2 Results and Discussion**

The traditional poverty line and its variants as defined in (2), (3) and (4) respectively are shown in Table (1).  $z_2$  is the highest followed by  $z_3$  and  $z_1$ . The corresponding number of the poor (q) from the poverty lines are shown in Table (2). The highest number of the poor came from  $z_2$  followed by the numbers of the poor from  $z_3$  and  $z_1$ . The estimates of the number of the poor based on the binomial test statistic for  $z_2$ ,  $z_3$  and  $z_1$  in Table (3) are 3000, 2269 and 1867 respectively.

$z_1$ (traditional poverty line)	$z_2$	$z_3$
<b>₦73989.9</b>	<b>₦110984.9</b>	<b>₦85906.95</b>

**Table 1: Estimates of the Poverty Lines**

$z_1$ (traditional poverty line)	$z_2$	$z_3$
<b>1867</b>	<b>3000</b>	<b>2269</b>

**Table 2: Estimates of the number of the poor**

As discussed under (1.3), the test statistic T of the binomial test is the number of the poor also. From Table (2), it is obvious that the test statistic, T, is as shown in Table (3) for the poverty lines  $z_1$ ,  $z_2$  and  $z_3$  respectively.

	<b>T</b>
$z_1$	<b>1867</b>
$z_2$	<b>3000</b>
$z_3$	<b>2269</b>

**Table 3: Estimates of the binomial test statistic for the poverty lines**

The large sample approximation to the binomial test as noted in (1.4) was used since the size (4536) of the data is indeed large. The mean, standard deviation and the test statistic values under the binomial test hypotheses are shown in Table (4). The means and standard deviations are the same for the three poverty lines used in this study. The reason for this similarity is because the sample size (n) and the hypothesized proportion of the poor ( $p^*$ ) are fixed. Varying the values of these binomial test parameters might be a suggestion for future research. The test statistic values for the large sample approximation reveal that the highest value came from  $z_2$  while the least came from  $z_1$ . The last column of Table (3.4) gives the critical value of the test statistic when  $\alpha = 0.05$ . The implication of this critical value is that the null hypothesis formulated in (2.6a) cannot be rejected since  $-Z^* < -Z_\alpha$ . For the hypothesis (2.6b), the null hypothesis will be rejected since  $-Z^* < -Z_\alpha$ . The implication of accepting and rejecting the null hypotheses in (2.6a) and (2.6b) is that  $p^* \geq 0.69$  and this validates the claim by Kale (2012). The results here further implies that poverty is on the increase in Nigeria and this corroborates Ajakaiye et al (2016, p.218) who noted that “ Nigeria, no doubt, typifies a country that has had rapid economic growth but worsening poverty”. This study, by this finding, correlates with IndexMundi’s (<http://goo.gl/bmKYvS>) estimate of the number of people living below the poverty line which was estimated as 70%.

	<b>Mean <math>E_o(T)</math></b>	<b>Standard Deviation <math>\sqrt{Var_o(T)}</math></b>	<b><math>Z^*</math></b>	
$z_1$	<b>3129.84</b>	<b>31.1488</b>	<b>-40.542</b>	<b><math>Z_\alpha = Z_{0.05} = 1.645</math></b>
$z_2$	<b>3129.84</b>	<b>31.1488</b>	<b>-4.168</b>	
$z_3$	<b>3129.84</b>	<b>31.1488</b>	<b>-27.636</b>	

**Table 4: Estimates of the mean, variance, standard deviation and  $Z^*$  under the null hypotheses of the binomial test for the poverty lines**



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### Conclusion

This study has attempted an approach of estimating the poverty incidence in Nigeria via the binomial test. The appeal of this approach is that it allows us to determine, statistically, whether the incidence of poverty has increased or decreased. This is needful because it allows the government to access the impact of poverty interventions. It is worthy of note that the binomial test is robust to different poverty lines and the finding of this study suggests that poverty incidence in Nigeria is on the rise in line with past studies. There is therefore the need for the government to accelerate sincere efforts in reducing the burden of poverty in the country.

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